Linear Bandits

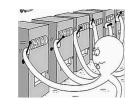
Sham M. Kakade and Wen Sun

Outline

- Recap
- Linear Bandits
 - Setting
 - LinUCB
 - An Optimal Regret Bound
- 3 Analysis
 - Regret Analysis
 - Confidence Analysis

Intro to MAB

Setting:



We have K many arms: $a_1, ..., a_K$

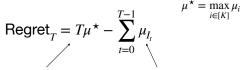
Each arm has a unknown reward distribution, i.e., $\nu_i \in \Delta([0,1])$, w/ mean $\mu_i = \mathbb{E}_{r \sim \nu}[r]$

Example: a_i has a Bernoulli distribution ν_i w/ mean $\mu_i := p$:

Every time we pull arm a_i , we observe an i.i.d reward $r = \begin{cases} 1 & \text{w/ prob } p \\ 0 & \text{w/ prob } 1-p \end{cases}$

Intro to MAB

More formally, we have the following learning objective:



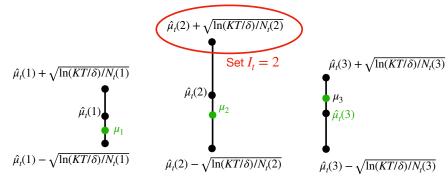
Total expected reward if we pulled best arm over T rounds

Total expected reward of the arms we pulled over T rounds

Goal: no-regret, i.e.,
$$\operatorname{Regret}_T/T \to 0$$
, as $T \to \infty$

UCB: Optimism in the face of Uncertainty

Given the confidence interval, we pick arm that has the highest Upper-Conf-Bound:



UCB Regret:

[Theorem (informal)] With high probability, UCB has the following regret:

$$\mathrm{Regret}_T = \, \widetilde{O} \left(\sqrt{KT} \, \right)$$

Generalization in RL

- (distribution free) Agnostic learning is not possible in RL: we showed that to get $O(\log |\Pi|)$ sample complexity we need either:
 - poly(|S|) samples OR
 - exp(H) samples.

in order to learn the best policy in some policy class.

upshot: we need stronger assumptions for RL analysis.

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- so the the conditional expectation of r_t is linear)
- Also, we have the noise sequence,

$$\eta_t = r_t - \mu^\star \cdot x_t$$

is i.i.d noise.

model due to Abe & Long '99

Our Objective

If $x_0, \dots x_{T-1}$ are our decisions, then our cumulative regret is

$$R_T = T\mu^{\star} \cdot x^{\star} - \sum_{t=0}^{T-1} \mu^{\star} \cdot x_t$$

where $x^* \in D$ is an optimal decision for μ^* , i.e.

$$\mathbf{X}^{\star} \in \operatorname{argmax}_{\mathbf{X} \in \mathbf{D}} \mu^{\star} \cdot \mathbf{X}$$

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The "Confidence Ball"

After t rounds, define our uncertainty region $BALL_t$: with center, $\widehat{\mu}_t$, and shape, Σ_t , using the λ -regularized least squares solution:

$$\widehat{\mu}_{t} = \arg\min_{\mu} \sum_{\tau=0}^{t-1} \|\mu \cdot x_{\tau} - r_{\tau}\|_{2}^{2} + \lambda \|\mu\|_{2}^{2}$$

$$= \Sigma_{t}^{-1} \sum_{\tau=0}^{t-1} r_{\tau} x_{\tau},$$

$$\Sigma_{t} = \lambda I + \sum_{\tau=0}^{t-1} x_{\tau} x_{\tau}^{\top}, \text{ with } \Sigma_{0} = \lambda I.$$

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Define the uncertainty region:

$$\mathsf{BALL}_t = \left\{ \mu \mid (\widehat{\mu}_t - \mu)^\top \Sigma_t (\widehat{\mu}_t - \mu) \leq \beta_t \right\},$$

where β_t is a parameter of the algorithm.

LinUCB (the algo)

- **1** Input: λ , β_t
- ② For t = 0, 1, ...
 - Execute

$$\mathbf{X}_t = \operatorname{argmax}_{\mathbf{X} \in D} \max_{\mu \in \mathsf{BALL}_t} \mu \cdot \mathbf{X}$$

and observe the reward r_t .

② Update $BALL_{t+1}$.

LinUCB Regret Bound

Sublinear regret: $R_T \leq O^*(d\sqrt{T})$ poly dependence on d, no dependence on the cardinality |D|.

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Theorem

Suppose: $|\mu^* \cdot x| \le 1$ and $||x|| \le B$ for all $x \in D$; that the noise is σ^2 sub-Gaussian; and that $||\mu^*|| \le W$. Set $\lambda = \sigma^2/W^2$ and

$$\beta_t := \sigma^2 \Big(2 + 4d \log \left(1 + \frac{tB^2 W^2}{d} \right) + 8 \log(4/\delta) \Big).$$

With probability greater than $1 - \delta$, that for all $T \ge 0$,

$$R_T \leq c\sigma\sqrt{T}\left(d\log\left(1+rac{TB^2W^2}{d\sigma^2}
ight) + \log(4/\delta)
ight)$$

where c is an absolute constant.

due to Dani, Hayes, K. '09

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Confidence

In establishing the upper bounds there are two main propositions from which the upper bounds follow. The first is in showing that the confidence region is valid.

Proposition

(Confidence) Let $\delta > 0$. We have that

$$\Pr(\forall t, \, \mu^* \in \mathsf{BALL}_t) \geq 1 - \delta.$$

Sum of Squares Regret Bound

Assuming the confidence event holds, the following controls on the growth of the regret.

Proposition

(Sum of Squares Regret Bound) Define:

$$\operatorname{regret}_t = \mu^{\star} \cdot \mathbf{X}^* - \mu^{\star} \cdot \mathbf{X}_t$$

Suppose $||x|| \le B$ for $x \in D$. Suppose β_t is increasing and larger than 1. Suppose $\mu^* \in \mathsf{BALL}_t$ for all t, then

$$\sum_{t=0}^{T-1} \operatorname{regret}_t^2 \le 4\beta_T d \log \left(1 + \frac{TB^2}{d\lambda}\right)$$

Completing the Proof

Proof:[Proof of Theorem 1] With the two previous Propositions, along with the Cauchy-Schwarz inequality, we have, with probability at least $1-\delta$,

$$R_T = \sum_{t=0}^{T-1} \operatorname{regret}_t \leq \sqrt{T \sum_{t=0}^{T-1} \operatorname{regret}_t^2} \leq \sqrt{4T\beta_T d \log \left(1 + \frac{TB^2}{d\lambda}\right)}.$$

The remainder of the proof follows from using our chosen value of β_T and algebraic manipulations.

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"Width" of Confidence Ball

Lemma

Let $x \in D$. If $\mu \in \mathsf{BALL}_t$ and $x \in D$. Then

$$|(\mu - \widehat{\mu}_t)^{\top} x| \le \sqrt{\beta_t x^{\top} \Sigma_t^{-1} x}$$

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$$|(\mu - \widehat{\mu}_t)^{\top} x| \le \sqrt{\beta_t x^{\top} \Sigma_t^{-1} x}$$

Proof: By Cauchy-Schwarz, we have:

$$\begin{aligned} &|(\mu - \widehat{\mu}_t)^\top x| = |(\mu - \widehat{\mu}_t)^\top \Sigma_t^{1/2} \Sigma_t^{-1/2} x| = |(\Sigma_t^{1/2} (\mu - \widehat{\mu}_t))^\top \Sigma_t^{-1/2} x| \\ &\leq \|\Sigma_t^{1/2} (\mu - \widehat{\mu}_t)\| \|\Sigma_t^{-1/2} x\| = \|\Sigma_t^{1/2} (\mu - \widehat{\mu}_t)\| \sqrt{x^\top \Sigma_t^{-1} x} \leq \sqrt{\beta_t x^\top \Sigma_t^{-1} x} \end{aligned}$$

where the last inequality holds since $\mu \in \mathsf{BALL}_t$.

Instantaneous Regret Lemma

Define

$$w_t := \sqrt{x_t^{\top} \Sigma_t^{-1} x_t}$$

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Lemma

Fix
$$t \leq T$$
. If $\mu^{\star} \in \mathsf{BALL}_t$, then

$$\operatorname{regret}_t \leq 2 \min \left(\sqrt{\beta_t} w_t, 1 \right) \leq 2 \sqrt{\beta_T} \min \left(w_t, 1 \right)$$

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Lemma

Fix $t \leq T$. If $\mu^* \in \mathsf{BALL}_t$, then

$$\operatorname{regret}_t \leq 2 \min \left(\sqrt{\beta_t} w_t, 1 \right) \leq 2 \sqrt{\beta_T} \min \left(w_t, 1 \right)$$

Proof: Let $\widetilde{\mu} \in \mathsf{BALL}_t$ denote the vector which maximizes the dot product $\widetilde{\mu}^\top x_t$. By choice of x_t , we have

$$\widetilde{\boldsymbol{\mu}}^{\top} \mathbf{X}_t = \max_{\boldsymbol{\mu} \in \mathsf{BALL}_t} \boldsymbol{\mu}^{\top} \mathbf{X}_t = \max_{\mathbf{X} \in D} \max_{\boldsymbol{\mu} \in \mathsf{BALL}_t} \boldsymbol{\mu}^{\top} \mathbf{X} \geq (\boldsymbol{\mu}^{\star})^{\top} \mathbf{X}^{*},$$

where the inequality used the hypothesis $\mu^{\star} \in \mathsf{BALL}_t$. Hence,

$$\operatorname{regret}_{t} = (\mu^{\star})^{\top} X^{*} - (\mu^{\star})^{\top} X_{t} \leq (\widetilde{\mu} - \mu^{\star})^{\top} X_{t}$$
$$= (\widetilde{\mu} - \widehat{\mu}_{t})^{\top} X_{t} + (\widehat{\mu}_{t} - \mu^{\star})^{\top} X_{t} \leq 2\sqrt{\beta_{t}} w_{t}$$

The next two lemmas give us 'geometric' potential function argument, where can bound the sum of widths independently of the choices made by the algorithm.

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Lemma

We have:

$$\det \Sigma_T = \det \Sigma_0 \prod_{t=0}^{r-1} (1 + w_t^2).$$

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Lemma

We have:

$$\det \Sigma_T = \det \Sigma_0 \prod_{t=0}^{r-1} (1 + w_t^2).$$

Proof: By the definition of Σ_{t+1} , we have

$$\begin{split} \det \Sigma_{t+1} &= \det (\Sigma_t + x_t x_t^\top) = \det (\Sigma_t^{1/2} (I + \Sigma_t^{-1/2} x_t x_t^\top \Sigma_t^{-1/2}) \Sigma_t^{1/2}) \\ &= \det (\Sigma_t) \det (I + \Sigma_t^{-1/2} x_t (\Sigma_t^{-1/2} x_t)^\top) = \det (\Sigma_t) \det (I + v_t v_t^\top), \end{split}$$

where $v_t := \sum_{t=1}^{-1/2} x_t$. Now observe that $v_t^{\top} v_t = w_t^2$ and ...

Lemma

For any sequence $x_0, \dots x_{T-1}$ such that, for t < T, $||x_t||_2 \le B$, we have:

$$\log\left(\det\Sigma_{T-1}/\det\Sigma_0\right) = \log\det\left(I + \frac{1}{\lambda}\sum_{t=0}^{T-1}x_tx_t^\top\right) \le d\log\left(1 + \frac{TB^2}{d\lambda}\right).$$

Lemma

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Proof: Denote the eigenvalues of $\sum_{t=0}^{T-1} x_t x_t^{\top}$ as $\sigma_1, \dots, \sigma_d$, and note:

$$\sum_{i=1}^{d} \sigma_i = \text{Trace}\Big(\sum_{t=0}^{T-1} x_t x_t^{\top}\Big) = \sum_{t=0}^{T-1} \|x_t\|^2 \le TB^2.$$

Using the AM-GM inequality,

$$\log \det \left(I + \frac{1}{\lambda} \sum_{t=0}^{T-1} x_t x_t^{\top}\right) = \log \left(\prod_{i=1}^{d} (1 + \sigma_i/\lambda)\right)$$

$$= d \log \left(\prod_{i=1}^{d} \left(1 + \sigma_i/\lambda\right)\right)^{1/d} \le d \log \left(\frac{1}{d} \sum_{i=1}^{d} (1 + \sigma_i/\lambda)\right) \le d \log \left(1 + \frac{TB^2}{d\lambda}\right)$$

Proving "sum of squares regret" Proposition

Proof:[Proof of Proposition 3] Assume $\mu^* \in BALL_t$ for all t. We have:

$$\begin{split} &\sum_{t=0}^{T-1} \mathrm{regret}_t^2 \leq \sum_{t=0}^{T-1} 4\beta_t \min(w_t^2, 1) \leq 4\beta_T \sum_{t=0}^{T-1} \min(w_t^2, 1) \\ &\leq 8\beta_T \sum_{t=0}^{T-1} \ln(1 + w_t^2) \leq 8\beta_T \log\left(\det \Sigma_{T-1} / \det \Sigma_0\right) \\ &= 8\beta_T d \log\left(1 + \frac{TB^2}{d\lambda}\right) \end{split}$$

where the first inequality follow from by Lemma 5; the second from that β_t is an increasing function of t; the third uses that for $0 \le y \le 1$, $\ln(1+y) \ge y/2$; the final two inequalities follow by Lemmas 6 and 7.

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Confidence [Proof of Proposition 2]

Proof: Since $r_{\tau} = x_{\tau} \cdot \mu^{\star} + \eta_{\tau}$, we have:

$$\widehat{\mu}_{t} - \mu^{*} = \sum_{t=0}^{t-1} \sum_{\tau=0}^{t-1} r_{\tau} x_{\tau} - \mu^{*} = \sum_{t=0}^{t-1} \sum_{\tau=0}^{t-1} x_{\tau} (x_{\tau} \cdot \mu^{*} + \eta_{\tau}) - \mu^{*}$$

$$= \sum_{t=0}^{t-1} \left(\sum_{\tau=0}^{t-1} x_{\tau} (x_{\tau})^{\top} \right) \mu^{*} - \mu^{*} + \sum_{t=0}^{t-1} \sum_{\tau=0}^{t-1} \eta_{\tau} x_{\tau}$$

$$= \lambda \sum_{t=0}^{t-1} \mu^{*} + \sum_{t=0}^{t-1} \eta_{\tau} x_{\tau}$$

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$$= \lambda \sum_{t=0}^{t-1} \mu^{*} + \sum_{t=0}^{t-1} \eta_{\tau} x_{\tau}$$

By the triangle inequality,

$$\sqrt{(\widehat{\mu}_t - \mu^*)^\top \Sigma_t (\widehat{\mu}_t - \mu^*)} \leq \left\| \lambda \Sigma_t^{-1/2} \mu^* \right\| + \left\| \Sigma_t^{-1/2} \sum_{\tau=0}^{t-1} \eta_\tau X_\tau \right\| \\
\leq \sqrt{\lambda} \|\mu^*\| + ??.$$

How can we bound "??" To be continued...

S. M. Kakade RL 20/22

Self-Normalizing Sum

Lemma (Self-Normalized Bound for Vector-Valued Martingales)

(Abassi et. al '11) Suppose $\{\varepsilon_i\}_{i=1}^{\infty}$ are mean zero random variables (can be generalized to martingales), and ε_i is bounded by σ . Let $\{X_i\}_{i=1}^{\infty}$ be a stochastic process. Define $\Sigma_t = \Sigma_0 + \sum_{i=1}^t X_i X_i^{\top}$. With probability at least $1 - \delta$, we have for all $t \geq 1$:

$$\left\| \sum_{i=1}^t X_i \varepsilon_i \right\|_{\Sigma_{-}^{-1}}^2 \leq \sigma^2 \log \left(\frac{\det(\Sigma_t) \det(\Sigma_0)^{-1}}{\delta^2} \right).$$

Continued... [Proof of Proposition 2]

Proof:

$$\begin{split} (\widehat{\mu}_t - \mu^{\star})^{\top} \Sigma_t (\widehat{\mu}_t - \mu^{\star}) &\leq \left\| \lambda \Sigma_t^{-1/2} \mu^{\star} \right\| + \left\| \Sigma_t^{-1/2} \sum_{\tau=0}^{t-1} \eta_{\tau} X_{\tau} \right\| \\ &\leq \sqrt{\lambda} \|\mu^{\star}\| + \sqrt{2\sigma^2 \log \left(\det(\Sigma_t) \det(\Sigma^0)^{-1} / \delta_t \right)}. \end{split}$$

We seek to lower bound $\Pr(\forall t, \mu^* \in \mathsf{BALL}_t)$. Assign failure probability $\delta_t = (3/\pi^2)/t^2$ for the t-th event, which gives us:

$$1 - \Pr(\forall t, \, \mu^* \in \mathsf{BALL}_t) = \Pr(\exists t, \, \mu^* \notin \mathsf{BALL}_t) \le \sum_{t=1}^{\infty} \Pr(\mu^* \notin \mathsf{BALL}_t)$$
$$< \sum_{t=1}^{\infty} (1/t^2)(3/\pi^2) = 1/2.$$

This along with Lemma 7 completes the proof.